

# 1 Age-Related Differences in Processing Unconventional Text 2 Formats

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## 12 **ABSTRACT**

13

14 Typographic formats influence reading efficiency; however, knowledge remains limited regarding how  
15 these effects change across the lifespan, especially for orthographic distortions in digital environments. This  
16 study examines how conventional formats (lowercase and uppercase) and unconventional formats (mixed-  
17 case and LEET) affect reading times and the integration of meaning while reading five-word phrases. Three  
18 hundred and three adults (18–84 years) read short sentences (five words) presented in the four formats,  
19 while reading times and memory accuracy were recorded. The results showed a graded cost pattern:  
20 conventional formats yielded the fastest reading times, mixed-case imposed moderate costs, and LEET  
21 produced the greatest slowdown and a slight reduction in accuracy. Moreover, a significant interaction  
22 between format and age was observed: although reading slowed with age in all formats, this effect was  
23 especially pronounced for LEET. These findings suggest that extreme orthographic distortions increase  
24 perceptual and pre-lexical demands, revealing limits in reading adaptation associated with aging.

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26 **Keywords:** Typographic formats; Reading efficiency; Orthographic distortions; LEET; Lifespan; Sentence  
27 reading

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29

## 30 **Introduction**

31 Studies in visual word recognition investigate how readers perceive, identify, and decode printed words,  
32 from the extraction of visual features through letter recognition to accessing meaning and pronunciation  
33 within the language system (Snowling, Hulme & Nation, 2022). This line of research aims to clarify the  
34 stepwise mechanisms underlying reading and written language comprehension, with an emphasis on how  
35 orthographic information activates linguistic representations. Among the multiple variables that modulate  
36 the reading process, recent literature has established that the physical properties of written words and their  
37 constituent letters (e.g., spacing, font width, orthographic format) influence the early stages of visual word  
38 recognition (Minakata & Beier, 2021; Scaltritti et al., 2019; Staub, 2020).

39 A particularly robust demonstration of orthographic format effects is the lowercase superiority effect,  
40 whereby, compared to words in uppercase, words in lowercase yield shorter response times and shorter  
41 fixation times (Tinker and Paterson, 1928; see also Perea, Fernández-López & Marcet, 2020; Perea, Rosa  
42 & Marcet, 2017). Two complementary explanations have been proposed for the lowercase superiority  
43 effect. First, greater exposure to lowercase text compared to all-uppercase—used mainly for titles and  
44 acronyms—establishes it as the perceptual default (Perea et al., 2017). Second, lowercase letters feature  
45 distinctive ascenders ("k", "d") and descenders ("j", "g") that create unique visual shapes absent in uniform-  
46 height uppercase letters (Paterson and Tinker, 1940).

47 From a developmental perspective, letter case sensitivity emerges gradually as children acquire  
48 alphabet knowledge and begin formal reading instruction. Early in literacy acquisition, children typically  
49 learn to recognize and name uppercase letters first, with accurate lowercase recognition lagging behind  
50 and improving as reading experience increases (e.g., van de Walle de Ghelcke et al., 2021). At these early  
51 stages, children often show a preference for one case over the other, but this preference quickly gives way  
52 to an abstract notion of letter identity that generalizes across visual forms (Grainger et al., 2008; Grainger,  
53 2022). As this abstract letter identity becomes established, case information is relegated to a secondary  
54 role and the reader develops the ability to identify letters efficiently regardless of whether they appear in  
55 uppercase or lowercase.

56 In adult readers, case effects become much smaller but do not disappear entirely, with behavioural  
57 studies showing a subtle yet reliable advantage for lowercase words and word sequences in accuracy and/or  
58 speed (e.g., Perea & Rosa, 2002; Vergara-Martínez et al., 2020). Recent work shows that the lowercase  
59 advantage reflects more efficient early visuo-orthographic processing and case-independent letter access  
60 (Fournet et al., 2022).

61 However, some formats disrupt letter-shape familiarity even more strongly than uniform uppercase,

62 because they alter not only letter size but also the internal pattern of cases within a word. Among these,  
63 the alternation of uppercase and lowercase letters within a word, namely, mixed-case, e.g., ‘LOTTERY’  
64 spelled as ‘LoTtErY’ (Mayall et al., 1997) represents a paradigmatic example. This alternation occurs  
65 naturally in many languages, as in proper names or sentence beginnings, and, in some languages, common  
66 nouns, but it does not occur in multiple instances within a string.

67 Perea et al. (2015) provided a seminal analysis of mixed-case processing, demonstrating that this format  
68 disrupts word recognition by impairing letter grouping and case uniformity. Although subsequent studies  
69 remain limited, the existing evidence consistently shows that mixed-case imposes greater processing  
70 demands than conventional formats, slowing reading and hindering efficient word recognition (Mayall et  
71 al., 2001; McClelland, 1977; Poulton, 1969). This format particularly disrupts orthographic processing  
72 except for proper names, where the initial capital can facilitate recognition (Peressotti et al., 2003). This  
73 disruption has been said to arise from breaking case uniformity and impairing letter grouping due to  
74 varying heights (Arditi & Cho, 2007; Mayall et al., 1997; Mayall & Humphreys, 1996).

75 While mixed-case alternations constitute a clear deviation from conventional word forms, they still  
76 preserve the original orthographic identity of each letter. Other forms of visual distortion go a step further  
77 by systematically replacing one element in the written word with another —either due to strong physical  
78 overlap between characters (e.g., o–O) or phonological similarity (e.g., φ–f, which map onto /f/)— thus  
79 altering the graphical identity of the written word more radically. A clear example is Greeklish, which can  
80 be considered a form of transliteration, although it often lacks a fixed one-to-one correspondence between  
81 source and target characters. In Greeklish, Greek words are written using Latin characters that are  
82 phonologically similar to the original letters (e.g., Greek *λοτάρια* ‘lottery’ rendered as *lotaría* or *lotaria*).  
83 This informal way of writing emerged in digital communication as a practical solution to software and  
84 hardware limitations on Greek-script input, and it remains frequent in online and messaging contexts  
85 (Chalamandaris et al., 2006). In this case, the alphabet changes, but phonological similarity between  
86 characters supports word recognition (Dimitropoulou et al., 2011). While Greeklish primarily reflects a  
87 functional adaptation to technological constraints, other formats, such as LEET, employ similar  
88 substitution principles for strategic purposes, including bypassing content moderation filters on digital  
89 platforms.

90 LEET (or L33T) represents a more extreme case within digital environments: letters are replaced by  
91 orthographically similar digits or symbols (e.g., ‘lottery’ → ‘L0TT3RY’; Perea et al., 2008; Grabbe, 2016).  
92 This substitution adds another layer of complexity to reading comprehension, as LEET disrupts word  
93 recognition by altering the letters rather than their format, thereby challenging the abstract letter identity

94 tolerance that normally facilitates efficient word processing (Fournet et al., 2022). The users of platforms  
95 like Twitch and online games utilize LEET as a strategy to bypass moderation filters and to avoid the  
96 detection of censoring systems in online platforms, exploiting the fact that human readers can still recover  
97 the intended word while early filtering systems could not (Märtens, et al., 2015).

98 Given its peculiar orthographic characteristics, LEET has attracted attention in experimental studies  
99 exploring how altered visual word forms are processed during reading. For example, Perea et al. (2008)  
100 showed in a masked priming lexical decision task that visually similar LEET primes (e.g., L0TT3RY–  
101 LOTTERY) produce priming effects comparable to identity primes (e.g., LOTTERY–LOTTERY),  
102 relative to control primes (e.g., L6TT2R7–LOTTERY). These findings suggest that the digit-to-letter  
103 conversion process in LEET depends on significant visual resemblance between the digit and the letter it  
104 substitutes. This resemblance allows digits to function as letter instances and activate lexical  
105 representations (Carreiras, Duñabeitia & Perea, 2007; Kinoshita et al., 2014; Perea et al., 2008), likely  
106 because readers are used to assimilating letter identities across a variety of fonts and handwriting styles  
107 (see Manso de Zuniga et al., 1991). As argued by Molinaro et al. (2010), when embedded in letter strings,  
108 visually similar digits are readily regularized and encoded in a letter-like manner, whereas the reverse  
109 process (letters embedded in digit strings) does not occur to the same extent, indicating an asymmetry  
110 between letter and digit processing systems.

111 Building on these perceptual mechanisms, it is also important to consider characteristics of the reader  
112 that may modulate how such formats are processed. Beyond these perceptual constraints, an understudied  
113 factor in the current literature on LEET and mixed-case, and even standard case formats, is the modulatory  
114 role of age. There are at least two ways in which age and age-related factors might be relevant for how  
115 readers process text presented in these alternative formats: age may influence reading either as a proxy for  
116 experience or through broader developmental and aging-related changes in reading and orthographic  
117 processing (Perea et al., 2008). Understanding these age-related differences requires considering how  
118 reading mechanisms evolve across the lifespan. Developmentally, young readers rely more on serial, fine-  
119 grained processing of letter identities and positions, gradually transitioning to more parallel and efficient  
120 lexical access as they gain reading experience (Acha & Perea, 2008). In adulthood, however, aging leads  
121 to declines in early visual stages and orthographic processing, causing older readers to depend increasingly  
122 on lexical-level information, with a reduced impact of lower-level orthographic detail (MacKay & Abrams,  
123 1998; Spieler & Balota, 2000; Froehlich et al., 2016). Older adults show deficits on tasks that depend  
124 heavily on high-spatial-frequency visual information (e.g., Coyne, 1981; Thapar et al., 2004), and  
125 performance in tasks such as masked letter discrimination declines substantially with age. Conversely,

126 although lexical-decision tasks often show apparent aging effects, more integrative analyses suggest that  
127 these effects may reflect peripheral encoding or general decision processes rather than deficits in lexical-  
128 semantic processing per se. Three examples are particularly relevant. First, diffusion-model analyses  
129 indicate that older adults' slower lexical-decision responses can be attributed partly to more conservative  
130 response criteria, rather than to poorer lexical evidence accumulation (Ratcliff et al., 2004). Second,  
131 entropy-based measures that integrate response time and accuracy suggest age-related decrements in  
132 peripheral encoding, but not necessarily in central semantic processing (Allen et al., 2004). Third, slope-  
133 and-intercept analyses point to a similar conclusion: age differences in semantic-memory tasks are  
134 expressed primarily as intercept shifts, whereas episodic-memory tasks show both intercept and slope  
135 changes, consistent with stronger age-related effects on central processing in episodic than semantic tasks  
136 (Allen et al., 2002).

137 Morphological processing is also fully preserved in advanced age (Duñabeitia et al., 2009). Normative  
138 studies also suggest that older adults often show equal or even superior vocabulary knowledge relative to  
139 younger adults, as reflected in larger lexical repertoires (Keuleers et al., 2015; Aguasvivas et al., 2020).  
140 Moreover, individual differences in vocabulary and print exposure can mitigate some age-related deficits,  
141 enabling more efficient lexical processing despite the general effects of aging (Cohen-Shikora & Balota,  
142 2016). Consistent with this view, older adults show greater impairments in peripheral letter encoding for  
143 visually demanding formats like mixed-case and unfamiliar fonts, particularly in terms of speed and  
144 efficiency (e.g., Allen & Madden, 1989; Allen et al., 1993; Allen et al., 2002, 2011; Mayall, 2001, 2002).  
145 While these studies indicate that aging exacerbates difficulties under nonstandard visual conditions, how  
146 specific atypical formats are processed across the adult lifespan remains underexplored, particularly for  
147 LEET and for direct contrasts between lowercase and uppercase word recognition, as prior work has  
148 typically relied on age-homogeneous samples and limited case manipulations. In addition, most studies  
149 have been done with single word lexical decision (with masked priming sometimes), that do not require  
150 any integration of meaning across words.

151 The present study fills these gaps by testing adults with ages ranging from 18 to 84 years and examining  
152 how integration of phrase meanings in conventional (lowercase, uppercase) and unconventional (mixed-  
153 case, LEET) text formats vary across the adult lifespan. Specifically, it builds on prior research that  
154 examined these formats individually to trace how case and format effects change over the course of adult  
155 life. Our study seeks to determine whether a stable preference for one case emerges or shifts with age,  
156 examining whether age-related changes in analytic/holistic processing balance and letter encoding  
157 efficiency modulate the disruptions caused by atypical orthography.

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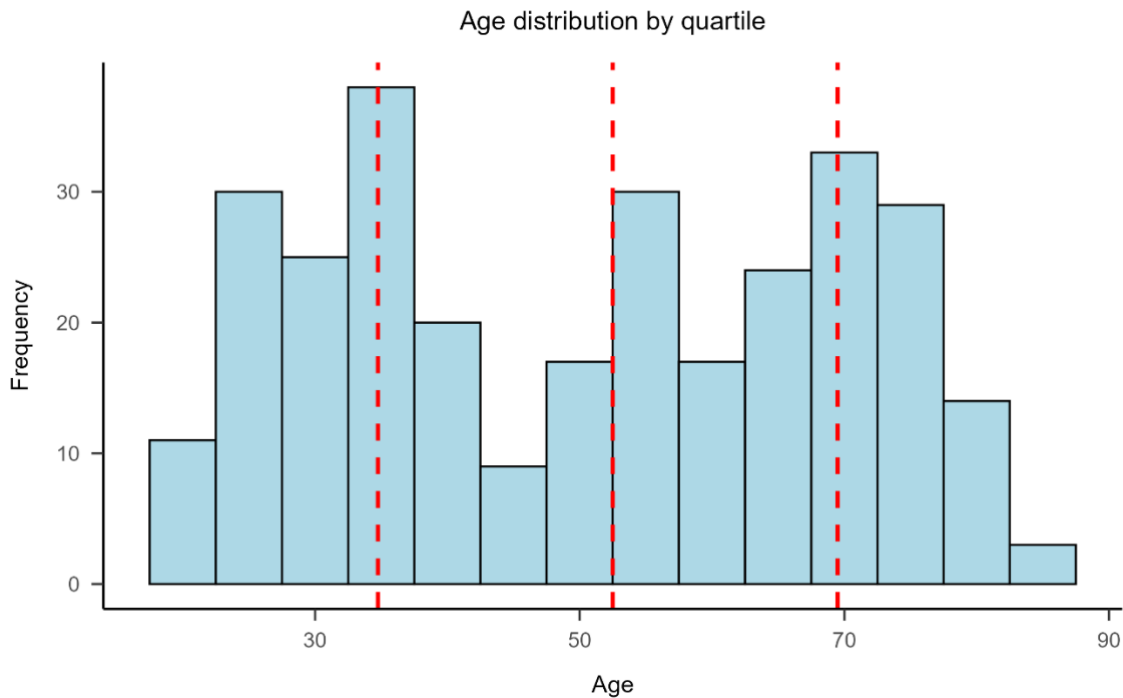
160 **Methods**

161 **Participants**

162 The sample consisted of 303 native English speakers (55% women, 44.7% men, 0.3% other), aged 18-  
163 84 years, recruited via Prolific Academic (Palan & Schitter, 2018). Figure 1 shows the age distribution of  
164 participants (mean = 50.81 years, SD = 18.67, range 18-84). Participants were divided into four age-based  
165 quartiles of equal size ( $n = 75$  each). The mean ages for Q1, Q2, Q3, and Q4 were 27.1, 41.3, 60.2, and  
166 74.6 years, respectively, with corresponding age ranges of 18–34, 35–52, 53–69, and 71–84 years.  
167 Participants provided written informed consent for their involvement in accordance with the Declaration  
168 of Helsinki. The study protocols were approved by the Research Ethics Committee at Nebrija University  
169 (UNNE-2022-0017). Participants were recruited using Prolific’s age and language-screening filters.  
170 Eligibility was restricted to native English speakers, and recruitment was designed to sample broadly  
171 across the adult lifespan. We did not administer a formal neuropsychological screening instrument for mild  
172 cognitive impairment; consequently, we cannot rule out the possibility that some older participants had  
173 undiagnosed cognitive impairment. The sample should therefore be interpreted as a broadly recruited adult  
174 lifespan sample rather than as a neuropsychologically screened healthy-aging cohort.

175

176 **Figure 1.** Distribution of participants’ ages across quartiles (quartile boundaries indicated by dotted red  
177 lines).



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180 **Materials**

181 The stimuli consisted of 27 unique English sentences presented in four orthographic formats (lowercase,  
 182 uppercase, mixed-case and LEET), for a total of 108 sentence stimuli. The 27 unique sentence stimuli  
 183 resulted from creating all possible combinations of three colours (*green, orange* and *white*), three numbers  
 184 (*five, eight* and *nine*) and three animals (*rabbit, elephant* and *tiger*), after which each sentence was rendered  
 185 in the four orthographic formats. The LEET numbers employed were: A = 4, E = 3 and I = 1. Sentences  
 186 followed the structure ‘There are [Number] [Colour] [Animal]’ (lowercase: ‘There are eight orange  
 187 elephants’; uppercase: ‘THERE ARE EIGHT ORANGE ELEPHANTS’; mixed-case: ‘ThErE aRe EiGhT  
 188 oRaNgE eLePhAnTs’; LEET: ‘TH3R3 4R3 31GHT 0R4NG3 3L3PH4NTS’). The presentation order was  
 189 randomized for each participant. All participants read each of the 108 total sentences once. In addition to  
 190 the sentences, square icons representing the three animals, three numbers and three colours were built to  
 191 use in the memory check after each sentence (see Figure 2 below).

192 **Task and Procedure**

193 The experiment was conducted using Gorilla (Anwyl-Irvine et al., 2020), an online experiment builder  
 194 and testing platform that allowed for accuracy and reading times to be automatically recorded. All  
 195 sentences were displayed in Courier font, size 55. Participants completed the task within a maximum of  
 196 45 minutes, including a scheduled break halfway through the experiment.

197 The procedure began with an instruction screen that clearly explained the task. This was followed by a

198 short practice phase consisting of four trials—one for each typographic format—to familiarize participants  
199 with the procedure. Upon completing the practice, a screen informed participants that the experimental  
200 task was about to begin. A progress bar was displayed throughout the experiment to indicate how much of  
201 the task remained.

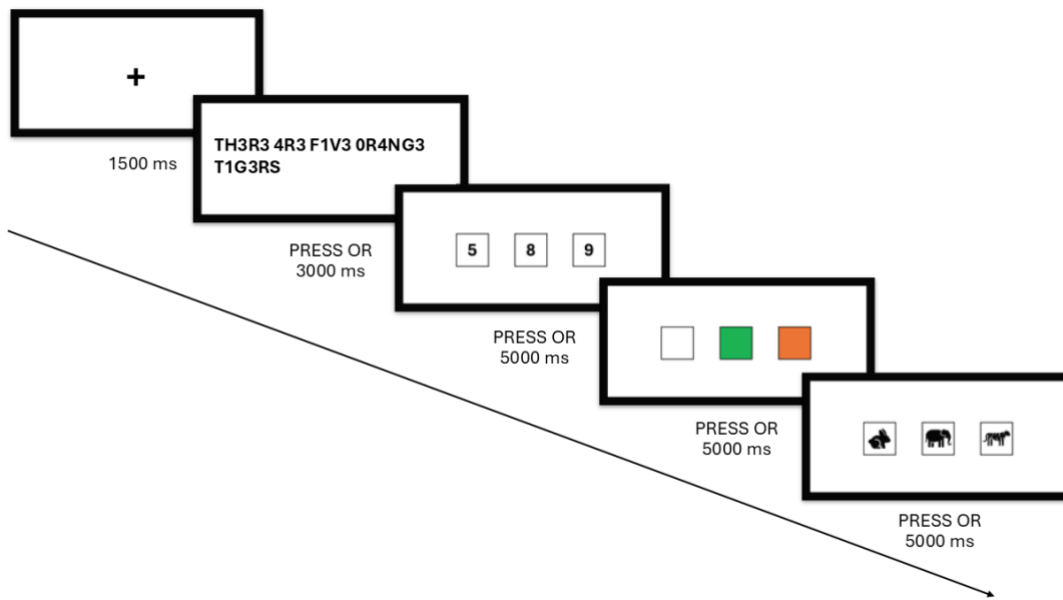
202 Each trial followed a fixed timeline. First, a fixation cross appeared at the centre of the screen for 1500  
203 ms to focus participants' attention and to standardize trial onset. Immediately afterward, the target sentence  
204 was presented along with the prompt "Press after reading." This prompt remained visible throughout the  
205 sentence display. Participants had a maximum of 3000 ms to read the sentence and press the spacebar. If  
206 they did not press the spacebar within this time, the sentence disappeared automatically and the experiment  
207 advanced to the next phase.

208 Following the reading phase, a memory check was conducted through three successive screens. In each  
209 screen, participants had to identify one of the key elements mentioned in the sentence: a number, a colour,  
210 or an animal. All the possible options for each element were displayed on each screen, and participants  
211 were required to select the one that corresponded to the sentence they had just read by clicking on the  
212 correct option. Each response screen had a maximum duration of 5000 ms; if no response was provided  
213 within this time, the trial was marked as a timeout. Two dependent variables were analysed: reading time  
214 and memory accuracy. Reading times were recorded as the latency between sentence onset and the  
215 participant's keypress. Memory accuracy was calculated as the percentage of correct selections across the  
216 three probes presented in each trial—i.e., a trial was considered correctly answered only if all three  
217 questions pertaining to that trial were answered correctly.

218

219 **Figure 2.** Examples of the task interface participants were exposed to. The first two screens depict the  
220 initial fixation plus reading sequence. The remaining three screens show the multiple-choice probes the  
221 participant performed to assess the meaning-integration of the phrase. *Note:* for readability, the font size  
222 in the figure has been increased, rendering two lines in the sentence screen. Sentences in the original  
223 experiment were displayed only on one line.

224



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## 227 **Results**

### 228 **Accuracy**

229 While our analyses focus predominantly on reading times, measuring accuracy was a necessary check  
 230 to ensure that participants were reading the sentences. Sentence memory was considered accurate only  
 231 when all three questions—regarding number, colour and animal— were answered correctly. From the  
 232 original sample of 303 participants, three were excluded due to low accuracy scores, < 67% (i.e., fewer  
 233 than 2/3 questions correct), which indicated a lack of attention to the task.

234 As expected, the proportion of accurate trials in the image selection task was high: for trials with  
 235 lowercase text it was 0.966 (SD = 0.055), for uppercase it was 0.960 (SD = 0.067), for mixed-case it was  
 236 0.966 (SD = 0.061), and for LEET it was 0.943 (SD = 0.084). This shows that, overall, participants were  
 237 able to read and understand the simple sentences regardless of format, and that they were paying attention  
 238 to the task.

239 Given the numerical differences across orthographic formats, we analyzed post-reading recognition  
 240 accuracy using a generalized linear mixed-effects model with a binomial family, fit with glmer from the R  
 241 package lme4 (Bates et al., 2015). Following the general principle of including the most complex random-  
 242 effects structure justified by the design, when possible (Barr et al., 2013), we fit the maximal model that  
 243 converged without singularity. The final model included fixed effects of Condition, Age, Trial, the  
 244 Condition × Age interaction, and the Trial × Age interaction, with random intercepts for Participant and  
 245 Item:

246  $glmer(Accuracy \sim Condition * Age\_centered +$   
247  $Trial\_centered * Age\_centered + (1 | Participant) +$   
248  $(1 | ItemN), family = binomial)$   
249

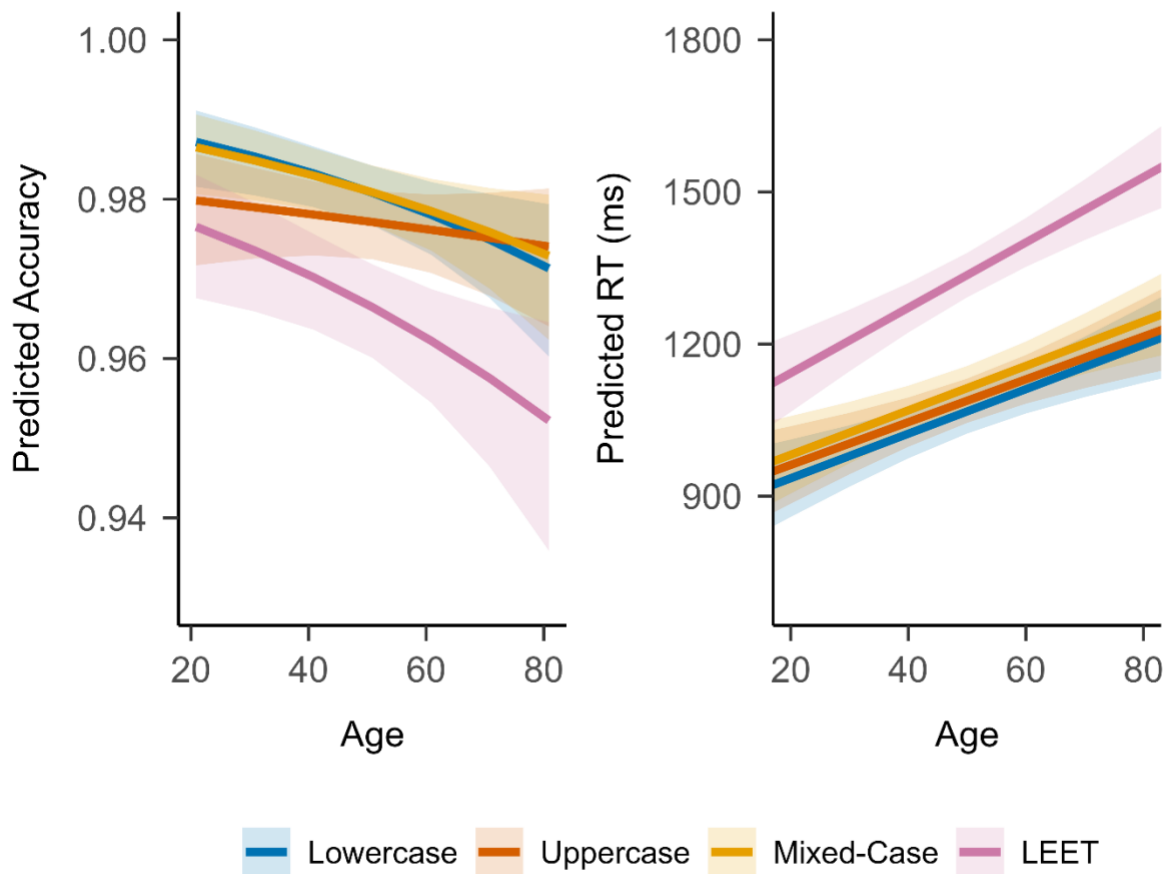
250 Accuracy data were analyzed at the probe-response level ( $n = 97,200$  observations). Descriptive  
251 statistics by orthographic format are reported in Table 1. LEET served as the reference level for the  
252 Condition factor. The model revealed a significant main effect of Age,  $\beta = -0.0115$ ,  $SE = 0.0046$ ,  $z =$   
253  $-2.53$ ,  $p = .011$ , indicating that post-reading recognition accuracy decreased with age. There was also a  
254 significant main effect of Trial,  $\beta = 0.0043$ ,  $SE = 0.0010$ ,  $z = 4.51$ ,  $p < .001$ , indicating that accuracy  
255 increased over the course of the experiment.

256 Accuracy was also affected by orthographic format. Compared with LEET, post-reading memory  
257 accuracy was significantly higher for lowercase sentences,  $\beta = 0.575$ ,  $SE = 0.096$ ,  $z = 6.02$ ,  $p < .001$ ;  
258 uppercase sentences,  $\beta = 0.391$ ,  $SE = 0.092$ ,  $z = 4.24$ ,  $p < .001$ ; and mixed-case sentences,  $\beta = 0.575$ ,  $SE =$   
259  $0.095$ ,  $z = 6.02$ ,  $p < .001$ . Thus, LEET yielded lower post-reading recognition accuracy than the other three  
260 orthographic formats.

261 The Condition  $\times$  Age interactions were not significant for lowercase relative to LEET,  $\beta =$   
262  $-0.0013$ ,  $SE = 0.0043$ ,  $z = -0.30$ ,  $p = .767$ , or for mixed-case relative to LEET,  $\beta = 0.0005$ ,  $SE =$   
263  $0.0043$ ,  $z = 0.12$ ,  $p = .909$ . The Uppercase  $\times$  Age interaction was marginal,  $\beta = 0.0078$ ,  $SE = 0.0041$ ,  $z =$   
264  $1.89$ ,  $p = .059$ , suggesting weak evidence that the age-related decline in accuracy may have been smaller  
265 for uppercase than for LEET. Importantly, the Trial  $\times$  Age interaction was significant,  $\beta = 0.00021$ ,  $SE =$   
266  $0.00005$ ,  $z = 4.15$ ,  $p < .001$ . A likelihood-ratio comparison confirmed that adding the Trial  $\times$  Age  
267 interaction improved model fit,  $\chi^2(1) = 17.50$ ,  $p < .001$ . This indicates that trial-related improvement in  
268 accuracy was stronger among older participants, consistent with practice, task familiarization, or  
269 adaptation to the restricted stimulus set over the course of the experiment.

270 To summarize, these results indicate that LEET imposed the greatest cost on post-reading recognition,  
271 that accuracy declined with age, and that participants—particularly older participants—showed evidence  
272 of improved accuracy as the experiment progressed.

273  
274 **Figure 3.** Predicted accuracy and reading time as a function of age, based on a generalized linear mixed-  
275 effects model described in the main text with a Format  $\times$  Age interaction. Coloured lines represent the  
276 estimated marginal means for each format, with shaded ribbons indicating 95% confidence intervals.



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282 **Table 1.** Proportion corrects and mean Reading Times (ms) by presentation format.

Format	Correct	SD	Mean RT	SD (by participant)
Lowercase	0.966	0.055	1067	323
Uppercase	0.960	0.067	1088	328

Mixed-Case	0.966	0.061	1117	340
LEET	0.943	0.084	1322	343

283 **Note.** SD for proportion correct computed across participants; SD for reading times computed by  
 284 participant.

285

286 **Reading Times**

287 Reading times associated with correct responses were analyzed. In addition, reading times shorter than  
 288 250ms (0.74% of all trials) and longer than 2500ms (2% of all trials) were eliminated.

289 Table 1 shows the mean and standard deviation (calculated across participants) for the four presentation  
 290 formats. Reading time was fastest for lowercase and slowest for LEET, with uppercase and mixed-case  
 291 clustering closer to lowercase (21 and 50 ms slower than lowercase, respectively, vs. 255 ms slower in the  
 292 case of LEET). These analyses were followed by exploratory data visualizations to examine the  
 293 distributional features of the observed effects via delta plots.

294 To analyze the effects of presentation format and participant age on reading times, we used a linear  
 295 mixed-effects model using the `lmer()` function with the *lmerTest* package in *R* (Kuznetsova et al., 2017).

296 
$$\text{Lmer}(\text{RT} \sim \text{Age} * \text{Format} + \text{TrialOrder} + (1 | \text{Participant}) + (1 | \text{Item}))$$

297 In this model, Format, Age and their interaction were specified as fixed effects, and random intercepts  
 298 were included for Participant and Item. LEET served as the reference level for the Format factor. Trial  
 299 Order (centered) was included in the model as a covariate to control for potential practice or fatigue effects  
 300 over the course of the session. We retained a random-intercepts-only structure for subjects and items  
 301 because more complex random-effects specifications led to singular fits, indicating that the data do not  
 302 support estimating additional variance components reliably.

303 Reaction Time data ( $n = 26,226$  observations) were analysed through a linear mixed-model analysis.  
 304 Results showed a significant main effect of Age (centred),  $F(1, 262.7) = 22.48, p < .001$ , demonstrating  
 305 that reaction times increased as a function of age. Furthermore, a significant main effect of Format was  
 306 found,  $F(3, 102.9) = 118.95, p < .001$ . Reading times in sentences displayed in LEET format were

307 significantly slower than in any other format (LEET vs. lowercase:  $\beta = -273.3$ ,  $SE = 16.38$ ,  $t(103.2) = -$   
308  $16.68$ ,  $p < 0.001$ ; LEET vs. uppercase:  $\beta = -250.6$ ,  $SE = 16.38$ ,  $t(103.3) = -15.30$ ,  $p < 0.001$ ; LEET vs.  
309 mixed-case:  $\beta = -225.72$ ,  $SE = 16.38$ ,  $t(103.3) = -13.77$ ,  $p < 0.001$ ). A significant main effect of Trial Order  
310 was also observed,  $F(1, 25,867.7) = 1627.16$ ,  $p < .001$ , showing that participants identified the words faster  
311 over the course of the experiment. Notably, a significant Age  $\times$  Format interaction was found,  $F(3,$   
312  $25,840.7) = 28.77$ ,  $p < .001$ , indicating that age-related slowing was more pronounced for the LEET format  
313 than for the other formats. In other words, while reading times increased with age across all formats, age-  
314 related slowing was particularly pronounced for the LEET format compared to the others. Figure 3  
315 illustrates this pattern, showing a clear separation between LEET and the remaining formats, as well as a  
316 general increase in reading times with age.

317 To examine condition-specific Age slopes, the model was reparametrized by changing the reference  
318 level of the Format factor (using treatment coding), so that each orthographic format's Age slope appeared  
319 directly as a fixed effect in the model output. Follow-up analyses showed that the slopes relating Age to  
320 reading time were significantly less steep for lowercase ( $\beta = -2.11$ ,  $SE = 0.28$ ,  $t(25840) = -7.46$ ,  $p < 0.001$ ),  
321 uppercase ( $\beta = -2.27$ ,  $SE = 0.28$ ,  $t(25840) = -7.99$ ,  $p < 0.001$ ), and mixed-case ( $\beta = -2.12$ ,  $SE = 0.28$ ,  
322  $t(25840) = -7.47$ ,  $p < 0.001$ ). To directly estimate the age slope for LEET, the model was reparametrized  
323 with Lowercase as the reference level, revealing a significant increase in reading time with age for LEET  
324 sentences ( $\beta = 2.11$ ,  $SE = 0.28$ ,  $t(25,840) = 7.46$ ,  $p < .001$ ).

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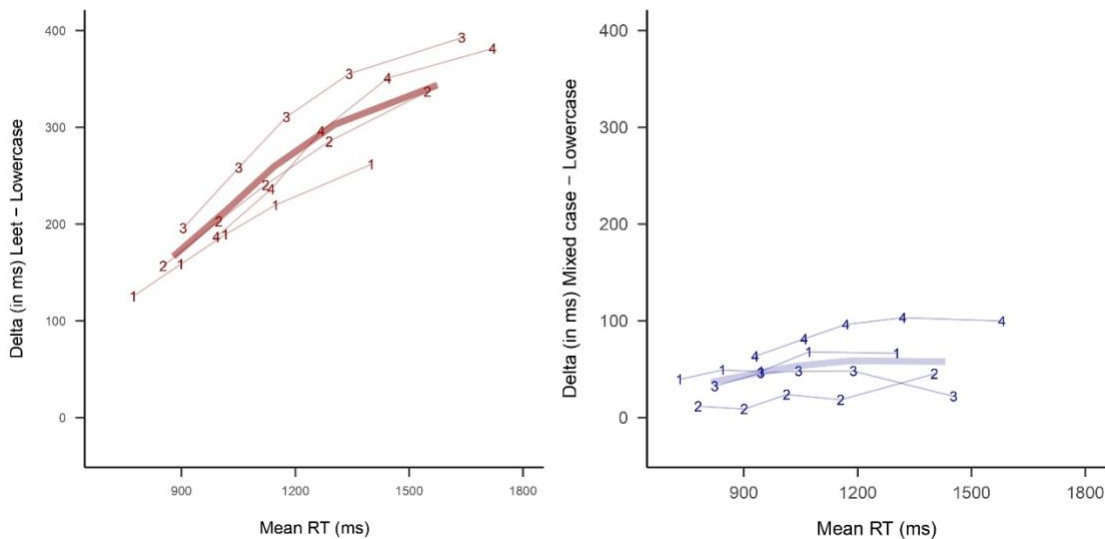
326 **Exploratory Data Analysis.** To complement the mixed-effects analyses of mean reading times, we  
327 conducted an exploratory data analysis (EDA) of the response-time distribution. EDA refers to descriptive  
328 and visualization-based analyses used to inspect patterns in the data beyond the primary confirmatory  
329 modeling. In the present case, this approach allowed us to examine whether orthographic-format costs  
330 were constant across the response-time distribution or became larger among slower responses. This  
331 distinction is theoretically important because these two possible patterns (flat delta plot: constant cost, vs  
332 sloped delta plots: larger effects for slower responses) indicate qualitatively different effects. We therefore  
333 constructed Vincentile-based delta plots using the 0.1, 0.3, 0.5, 0.7, and 0.9 quantiles of correct-response  
334 reading times. To construct these plots, we first computed each participant's reading-time quantiles  
335 separately for each orthographic condition. We then averaged the corresponding quantiles across  
336 participants to obtain group-level Vincentiles for each condition. For each pairwise comparison, the x-axis  
337 value at each quantile was computed as the mean of the two condition-specific Vincentiles being  
338 compared, whereas the y-axis value was computed as their difference. Thus, each point represents the

339 orthographic-format cost at a given location in the reading-time distribution. Lowercase sentences served  
340 as the baseline format, and deltas were computed for mixed-case, and LEET formats relative to this  
341 baseline (the delta plot for the lowercase-uppercase comparison is not shown for the sake of simplicity,  
342 but it is available in the OSF repository for the study).

343

344 **Figure 4.** Vincentile-based delta plots comparing reading times for LEET and mixed-case sentences  
345 relative to lowercase sentences. Delta values were computed at the 0.1, 0.3, 0.5, 0.7, and 0.9 quantiles of  
346 correct-response reading times. The x-axis represents the mean RT for the two conditions being compared  
347 at each quantile, and the y-axis represents the RT difference between conditions. Thick lines show the  
348 overall delta functions, and thinner lines show delta functions computed separately for age-based quartiles,  
349 labeled 1 through 4 from youngest to oldest. The LEET–lowercase contrast showed large positive deltas  
350 that increased at slower Vincentiles, particularly among older participants. In contrast, the mixed-case–  
351 lowercase contrast showed smaller and flatter deltas, indicating a more uniform cost across the RT  
352 distribution.

353



362 To examine whether these distributional patterns varied across ages for the mixed-case and LEET  
363 conditions, we divided participants into age-based quartiles and constructed delta plots separately for each  
364 quartile. Figure 4 shows the overall delta functions as thick lines, with quartile-specific delta functions  
365 overlaid and labeled 1 through 4, corresponding to increasing age. The LEET–lowercase contrast showed  
366 a strongly positive and upward-sloping delta function. This pattern was especially pronounced for the 3<sup>rd</sup>  
367 and 4<sup>th</sup> quartile ages, but is qualitatively similar to the 1<sup>st</sup> and 2<sup>nd</sup> quartiles. On the other hand, the flat delta

368 plot for the mixed case is higher for the oldest participants, but it does not appear to have a slope similar  
369 to that for the LEET-lowercase comparison.

370  
371 **Discussion**

372 The goal of this study was to examine how variations in typographical formats (lowercase, uppercase,  
373 mixed-case, LEET) affect reading times and memory for the read sentences across the adult lifespan (18-  
374 84 years), building on established effects like lowercase superiority (Perea et al., 2017, 2020) and case-  
375 mixing costs (Mayall et al., 1997). As noted in the introduction, it was expected that LEET would involve  
376 longer reading times, suggesting a higher cognitive load compared to conventional formats (lowercase,  
377 uppercase, mixed-case).

378 The results confirmed these predictions. Conventional formats yielded the fastest reading times and  
379 high accuracy, mixed-case imposed moderate costs, and LEET generated substantial disruptions in  
380 reading, accompanied by slightly reduced recognition. Although accuracy remained high across all  
381 formats, reading sentences in the LEET format resulted in reliably lower accuracy across age groups.

382 Crucially, analyses of reading times revealed that the disruptive effect of LEET increased with age,  
383 whereas age-related slowing was less pronounced for conventional formats, yielding a significant Format  
384 by Age interaction that would not be observable in age-homogeneous samples. These results are consistent  
385 with prior research documenting the lowercase superiority effect, whereby lowercase words are processed  
386 faster and more efficiently than uppercase ones (Perea et al., 2017; Vergara-Martínez et al., 2020; Perea et  
387 al., 2020). They also align with evidence that orthographic distortions, such as case mixing, disrupt reading  
388 fluency by breaking familiar visual regularities and increasing processing demands (Arditi & Cho, 2007;  
389 Mayall et al., 1997; Tinker, 1963).

390 Beyond mean effects, delta-plot analyses delineate the nature of these processing costs. Vincentile-  
391 based delta plots comparing Lowercase with LEET and mixed-case formats showed qualitatively different  
392 effects for LEET vs mixed-case. Notably, the lowercase-LEET contrast exhibited steeper slopes toward  
393 slower responses—particularly in older participants—suggesting that LEET engages processing  
394 mechanisms that place disproportionate demands on perceptual encoding and integration as reading  
395 unfolds. This pattern extends masked priming findings, in which visually similar LEET primes can activate  
396 lexical representations (Carreiras et al., 2007; Kinoshita et al., 2014), to naturalistic reading contexts where  
397 age-related constraints amplify the costs of orthographic distortion.

398 One of our goals was to explore whether age-related advantages due to increased exposure to reading  
399 with age might modulate the disruptions triggered by the atypical orthography in our experiment. We

400 arrive to a split-decision on this issue: The difference between LEET and mixed-case illuminate how  
401 distortions differentially tax visual word recognition mechanisms: while mixed-case preserves abstract  
402 letter identity access through standard alphabetic characters (Fournet et al., 2022), LEET imposes  
403 additional symbol-to-letter recoding demands, as evidenced by delta plots showing costs magnified for  
404 slower responses (Perea et al., 2008).

405 Age modulated these format effects in a manner consistent with prior evidence of heightened  
406 vulnerabilities in older adults' under visually demanding conditions, particularly at early perceptual and  
407 pre-lexical stages of processing (Allen et al., 2002, 2004; Mayall, 2001, 2002) rather than reflecting age-  
408 related changes in lexical representations per se (Perea et al., 2008). Specifically, disruption in the LEET  
409 format amplified with age while the disruption caused by mixed-case format showed comparatively  
410 smaller age-related costs and a flat delta plot pattern.

411 This Format  $\times$  Age interaction indicates that aging exacerbates the impact of orthographic irregularities  
412 that degrade visual regularities feeding into the lexical system, without implying a general decline in  
413 lexical processing. In this sense, LEET appears to place disproportionate demands on perceptual encoding  
414 and integration mechanisms that precede lexical access, thereby revealing age-related differences that are  
415 not apparent for visually canonical formats in age-homogeneous samples.

416 Taken together, these results suggest that the mechanisms underlying LEET processing differ  
417 qualitatively from those involved in case mixing. Whereas mixed-case disrupts global visual regularity,  
418 LEET requires an additional layer of symbol-to-letter recoding that particularly taxes older readers, whose  
419 analytic processing resources are more limited (Allen et al., 2011). Age effects were most pronounced for  
420 LEET: while younger readers tolerated LEET through flexible letter identity normalization (likely  
421 reflecting both digital exposure and preserved early visual processing), older adults showed  
422 disproportionate slowdowns and accuracy losses in the memory probes. This interpretation is further  
423 supported by delta plot analyses, which revealed that the cost of LEET increased disproportionately for  
424 slower responses, and aligns with lifespan shifts toward holistic, vocabulary-supported routes where  
425 compensatory lexical gains (Keuleers et al., 2015; Aguasvivas et al., 2020) preserve comprehension but  
426 cannot offset early perceptual demands (Spieler & Balota, 2000; Cohen-Shikora & Balota, 2016).  
427 Consequently, while the lowercase advantage and mixed-case disruption have been established before, our  
428 findings extend the literature by demonstrating how LEET's unique processing demands interact with age-  
429 related declines in peripheral orthographic encoding.

430 From a practical perspective, these findings underscore the challenges that unconventional formats such  
431 as LEET impose on reading efficiency and comprehension, particularly for older adults. Unlike children's

432 gradual abstraction of case-invariant letter identities across conventional formats (Grainger, 2022), adult  
433 aging reveals limits to this abstraction under extreme orthographic distortion: while developmental gains  
434 reflect experience-driven parallelization of skilled reading (Acha & Perea, 2008; Perea et al., 2017), aging  
435 reverses this efficiency for visually irregular formats like LEET. Although our data did not directly  
436 measure participants' prior exposure to LEET, the consistently slower reading times and reduced accuracy  
437 (disproportionate in older adults) suggest that familiarity alone may not be sufficient to compensate for  
438 LEET's demands on residual analytic resources beyond what compensatory lexical mechanisms can offset  
439 (Vergara-Martínez et al., 2020).

440 Naturally, this study contains some limitations. First, the experimental materials consisted of short,  
441 artificial sentences and limited LEET substitutions (A=4, E=3, I=1), constraining ecological validity since  
442 real LEET use involves diverse, context-specific substitutions encountered in chat messages and gaming  
443 environments. Second, participants' prior familiarity with LEET was not measured, even though previous  
444 exposure, and age as its proxy, could plausibly modulate both accuracy and reading times, limiting causal  
445 attribution between format per se and processing costs. Third, the restricted stimulus set involved repeated  
446 exposure to the same sentence constituents across trials, although no sentence was repeated within a given  
447 orthographic format. This repetition was useful for maintaining close comparability across conditions and  
448 for estimating distributional properties of reading times, but it may also have promoted learning or  
449 adaptation over the course of the experiment. In particular, participants may have become increasingly  
450 efficient at mapping recurring numbers, colors, and animals in unconventional formats such as mixed case  
451 or LEET onto their corresponding words.

452 Future work should therefore incorporate more ecological tasks, such as real chat messages or online  
453 conversations, to better approximate the contexts in which LEET is typically encountered. Moreover,  
454 combining behavioral measures with eye-tracking and EEG would provide a more precise characterization  
455 of the perceptual and lexical integration processes involved. In addition, a forthcoming study with Spanish  
456 participants residing in Morocco will explore these aspects in individuals regularly exposed to Darija, a  
457 dialect of Moroccan Arabic whose writing system incorporates the LEET format. Such a population offers  
458 a valuable opportunity to determine whether sustained exposure to LEET reduces its processing costs,  
459 thereby clarifying the extent to which familiarity and linguistic environment can mitigate the cognitive  
460 challenges posed by unconventional orthographic formats across development, adulthood, and aging.

461 In conclusion, this study demonstrates graded orthographic processing costs (conventional formats <  
462 mixed-case < LEET) that interact with age across the adult lifespan (18-84 years), confirming established  
463 effects like lowercase superiority and case-mixing disruption while revealing LEET's unique symbol-to-

464 letter recoding demands. By tracing format effects from young adulthood to advanced age, these findings  
465 integrate developmental, adult, and aging research, highlighting analytic processing decline as the critical  
466 vulnerability to digital orthographies. The current study not only contributes to the basic science of reading,  
467 but results advance equitable typography guidelines for education, digital design, and communication  
468 contexts that minimize age-related barriers across digital ecosystems.

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## **Competing interests**

The authors declare no competing interests.